



CIMAT

Centro de Investigación en Matemáticas, A.C.

High-frequency statistics for Levy processes: a Stein's method perspective

Arturo Jaramillo Gil, Chiara Amorino and Mark Podolskij

Centro de Investigación en Matemáticas (CIMAT)

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

How can we describe the fluctuation $F = S - \hat{S}$ of this approximation?

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

How can we describe the fluctuation $F = S - \hat{S}$ of this approximation?

Example:

- $\hat{S} = \mathbb{E}[S] =: \mu$.

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

How can we describe the fluctuation $F = S - \hat{S}$ of this approximation?

Example:

- $\hat{S} = \mathbb{E}[S] =: \mu$.
- Fluctuation is $F = S - \mu$.

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

How can we describe the fluctuation $F = S - \hat{S}$ of this approximation?

Example:

- $\hat{S} = \mathbb{E}[S] =: \mu$.
- Fluctuation is $F = S - \mu$.
- If S is the sum of n i.i.d. random variables with finite moments, then $F \stackrel{Law}{\approx} \sigma N$, with $N \sim \mathcal{N}(0, 1)$ and $\sigma^2 = \text{Var}[S]$.

Motivation

Let S be a random variable of interest that is not accessible, and let \hat{S} be a good approximation of S .

How can we describe the fluctuation $F = S - \hat{S}$ of this approximation?

Example:

- $\hat{S} = \mathbb{E}[S] =: \mu$.
- Fluctuation is $F = S - \mu$.
- If S is the sum of n i.i.d. random variables with finite moments, then $F \stackrel{Law}{\approx} \sigma N$, with $N \sim \mathcal{N}(0, 1)$ and $\sigma^2 = \text{Var}[S]$.
- Berry-Esseen theorem: $d_K(\frac{S-\mu}{\sigma}, N) \leq C/\sqrt{n}$.

General Objective

Understand the cumulative error when approximating a time series with a good predictor.

General Objective

Understand the cumulative error when approximating a time series with a good predictor.

Ingredients:

1. A non-accessible time series $\xi = \{\xi_k\}_{k \geq 1}$.

General Objective

Understand the cumulative error when approximating a time series with a good predictor.

Ingredients:

1. A non-accessible time series $\xi = \{\xi_k\}_{k \geq 1}$.
2. A very good estimator $\eta = \{\eta_k\}_{k \geq 1}$ of ξ .

General Objective

Understand the cumulative error when approximating a time series with a good predictor.

Ingredients:

1. A non-accessible time series $\xi = \{\xi_k\}_{k \geq 1}$.
2. A very good estimator $\eta = \{\eta_k\}_{k \geq 1}$ of ξ .

General Objective

Understand the cumulative error when approximating a time series with a good predictor.

Ingredients:

1. A non-accessible time series $\xi = \{\xi_k\}_{k \geq 1}$.
2. A very good estimator $\eta = \{\eta_k\}_{k \geq 1}$ of ξ .

Object of interest:

$$\text{Cumulative error} = (\xi_1 - \eta_1) + \cdots + (\xi_n - \eta_n)$$

True context of the problem

I observe a real-valued process $X = \{X_t\}_{t \geq 0}$ in stages:

True context of the problem

I observe a real-valued process $X = \{X_t\}_{t \geq 0}$ in stages: at stage i , I can observe $X_{1/n}, \dots, X_{(i-1)/n}$.

True context of the problem

I observe a real-valued process $X = \{X_t\}_{t \geq 0}$ in stages: at stage i , I can observe $X_{1/n}, \dots, X_{(i-1)/n}$.

Quantities of interest

At stage i , I would like to know

$$\xi_i := g\left(X_{\frac{i-1}{n}}, a_n \Delta_i X\right),$$

with $\Delta_i X := X_{\frac{i}{n}} - X_{\frac{i-1}{n}}$, $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ an appropriate function, and a_n an appropriate scaling.

True context of the problem

I observe a real-valued process $X = \{X_t\}_{t \geq 0}$ in stages: at stage i , I can observe $X_{1/n}, \dots, X_{(i-1)/n}$.

Quantities of interest

At stage i , I would like to know

$$\xi_i := g(X_{\frac{i-1}{n}}, a_n \Delta_i X),$$

with $\Delta_i X := X_{\frac{i}{n}} - X_{\frac{i-1}{n}}$, $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ an appropriate function, and a_n an appropriate scaling.

Observation mechanism

At stage i , I only have the information $\mathcal{F}_{i-1} := \sigma(X_{\frac{1}{n}}, \dots, X_{\frac{i-1}{n}})$.

True context of the problem

I observe a real-valued process $X = \{X_t\}_{t \geq 0}$ in stages: at stage i , I can observe $X_{1/n}, \dots, X_{(i-1)/n}$.

Quantities of interest

At stage i , I would like to know

$$\xi_i := g(X_{\frac{i-1}{n}}, a_n \Delta_i X),$$

with $\Delta_i X := X_{\frac{i}{n}} - X_{\frac{i-1}{n}}$, $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ an appropriate function, and a_n an appropriate scaling.

Observation mechanism

At stage i , I only have the information $\mathcal{F}_{i-1} := \sigma(X_{\frac{1}{n}}, \dots, X_{\frac{i-1}{n}})$.

Estimator:

$$\eta_i := \mathbb{E}[g(X_{\frac{i-1}{n}}, \Delta_i X) \mid \mathcal{F}_{i-1}].$$

Now indeed: the true context of the problem

Suppose that X starts at zero and has independent and stationary increments.

Now indeed: the true context of the problem

Suppose that X starts at zero and has independent and stationary increments.

For $t \geq 0$, we define the cumulative error on $[0, t]$ as

$$Z_n(t) := \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} (g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}]),$$

with $\mathcal{I}_{i,n} := a_n(X_{\frac{i}{n}} - X_{\frac{i-1}{n}}, \dots, X_{\frac{i+m}{n}} - X_{\frac{i+m-1}{n}})$, and $g : \mathbb{R}^{m+2} \rightarrow \mathbb{R}$ an appropriate function.

Now indeed: the true context of the problem

Suppose that X starts at zero and has independent and stationary increments.

For $t \geq 0$, we define the cumulative error on $[0, t]$ as

$$Z_n(t) := \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} (g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}]),$$

with $\mathcal{I}_{i,n} := a_n(X_{\frac{i}{n}} - X_{\frac{i-1}{n}}, \dots, X_{\frac{i+m}{n}} - X_{\frac{i+m-1}{n}})$, and $g : \mathbb{R}^{m+2} \rightarrow \mathbb{R}$ an appropriate function.

Assumptions

- The scaling a_n is such that $a_n X_{1/n}$ converges in law.
- There exists a constant $\alpha > 0$ such that $\mathbb{P}[X \geq s] \leq Cts^{-\alpha}$.

First problem of interest

Our **object of interest** is a process $Z = \{Z_n(t)\}_{t \geq 0}$, defined by

$$Z_n(t) = \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} (g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}]),$$

Key elements:

- Parameter α indicating non-integrability of X .
- Filter function g .

Questions of interest

- What is the limit of Z_n ? (if it exists)
- What is the rate of convergence?

Case where S is a r.v. and N is normal

The discrepancy between $F := (S - \mu)/\sigma$ and N is studied using expressions of the form

$$|\mathbb{E}[h(F) - h(N)]|. \quad (1)$$

Case where S is a r.v. and N is normal

The discrepancy between $F := (S - \mu)/\sigma$ and N is studied using expressions of the form

$$|\mathbb{E}[h(F) - h(N)]|. \quad (1)$$

Note that $F \stackrel{Law}{=} N$ if the following **property** holds: “(1) is zero for sufficiently many test functions h ”.

Case where S is a r.v. and N is normal

The discrepancy between $F := (S - \mu)/\sigma$ and N is studied using expressions of the form

$$|\mathbb{E}[h(F) - h(N)]|. \quad (1)$$

Note that $F \stackrel{Law}{=} N$ if the following **property** holds: “(1) is zero for sufficiently many test functions h ”.

Another **property** characterizing the normal distribution is...

Case where S is a r.v. and N is normal

The discrepancy between $F := (S - \mu)/\sigma$ and N is studied using expressions of the form

$$|\mathbb{E}[h(F) - h(N)]|. \quad (1)$$

Note that $F \stackrel{Law}{=} N$ if the following **property** holds: “(1) is zero for sufficiently many test functions h ”.

Another **property** characterizing the normal distribution is...

$$|\mathbb{E}[Nf'(N) - f''(N)]| = 0, \quad (2)$$

for sufficiently many functions f .

Case where S is a r.v. and N is normal

The discrepancy between $F := (S - \mu)/\sigma$ and N is studied using expressions of the form

$$|\mathbb{E}[h(F) - h(N)]|. \quad (1)$$

Note that $F \stackrel{Law}{=} N$ if the following **property** holds: “(1) is zero for sufficiently many test functions h ”.

Another **property** characterizing the normal distribution is...

$$|\mathbb{E}[Nf'(N) - f''(N)]| = 0, \quad (2)$$

for sufficiently many functions f .

Stein's heuristic

$$|\mathbb{E}[Ff'(F) - f''(F)]| \approx 0 \quad \Rightarrow \quad |\mathbb{E}[h(F) - h(N)]| \approx 0.$$

Formalization of Stein's heuristic

For a given function h , consider the equation

$$\mathbf{x} \cdot \nabla f(\mathbf{x}) - \text{Tr}[\text{Hess}[f](\mathbf{x})\Sigma] = h(\mathbf{x}) - \mathbb{E}[f(\mathbf{N})].$$

Formalization of Stein's heuristic

For a given function h , consider the equation

$$\mathbf{x} \cdot \nabla f(\mathbf{x}) - \text{Tr}[\text{Hess}[f](\mathbf{x})\Sigma] = h(\mathbf{x}) - \mathbb{E}[f(\mathbf{N})].$$

The solution is given by

$$f(\mathbf{x}) = \int_0^\infty (\mathbb{E}[h(\mathbf{N})] - \mathbb{E}_{\mathbf{x}}[h(\mathbf{Y}_t)]) dt,$$

where \mathbf{Y}_t is a Markov process starting at \mathbf{x} and converging to \mathbf{N} .

Formalization of Stein's heuristic

For a given function h , consider the equation

$$\mathbf{x} \cdot \nabla f(\mathbf{x}) - \text{Tr}[\text{Hess}[f](\mathbf{x})\Sigma] = h(\mathbf{x}) - \mathbb{E}[f(\mathbf{N})].$$

The solution is given by

$$f(\mathbf{x}) = \int_0^\infty (\mathbb{E}[h(\mathbf{N})] - \mathbb{E}_{\mathbf{x}}[h(\mathbf{Y}_t)]) dt,$$

where \mathbf{Y}_t is a Markov process starting at \mathbf{x} and converging to \mathbf{N} .

We obtain

$$\mathbb{E}[\mathbf{F} \cdot \nabla f(\mathbf{F}) - \text{Tr}[\text{Hess}[f](\mathbf{F})\Sigma]] = \mathbb{E}[h(\mathbf{F}) - f(\mathbf{N})].$$

Central issue of Stein's method

If \mathbf{F} and \mathbf{N} are multivariate and \mathbf{N} has covariance Σ , the quantity we need to control is

$$|\mathbb{E}[\mathbf{F} \cdot \nabla f_{\Sigma}(\mathbf{F}) - \text{Tr}[\text{Hess}[f_{\Sigma}](\mathbf{F})\Sigma]]|.$$

Central issue of Stein's method

If \mathbf{F} and \mathbf{N} are multivariate and \mathbf{N} has covariance Σ , the quantity we need to control is

$$|\mathbb{E}[\mathbf{F} \cdot \nabla f_{\Sigma}(\mathbf{F}) - \text{Tr}[\text{Hess}[f_{\Sigma}](\mathbf{F})\Sigma]]|.$$

Main challenge of the method:

How do we estimate $\mathbb{E}[\mathbf{F} \cdot \nabla f_{\Sigma}(\mathbf{F})]$?

Central issue of Stein's method

If \mathbf{F} and \mathbf{N} are multivariate and \mathbf{N} has covariance Σ , the quantity we need to control is

$$|\mathbb{E}[\mathbf{F} \cdot \nabla f_{\Sigma}(\mathbf{F}) - \text{Tr}[\text{Hess}[f_{\Sigma}](\mathbf{F})\Sigma]]|.$$

Main challenge of the method:

How do we estimate $\mathbb{E}[\mathbf{F} \cdot \nabla f_{\Sigma}(\mathbf{F})]$?

Typical approach: use the original ideas from Lindeberg's (or Stein's) method.

So what if $F = Z_n$? and Σ is random

The key computation is $\mathbb{E}[Z_n \cdot \nabla f_{\Sigma}(Z_n)]$.

So what if $F = Z_n$? and Σ is random

The key computation is $\mathbb{E}[Z_n \cdot \nabla f_{\Sigma}(Z_n)]$. In this case,

$$\begin{aligned} & \mathbb{E}[Z_n \cdot \nabla f_{\Sigma}(Z_n)] \\ &= \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} \mathbb{E}[(g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}])] \cdot \nabla f_{\Sigma}(Z_n) \end{aligned}$$

Key observation

So what if $F = Z_n$? and Σ is random

The key computation is $\mathbb{E}[Z_n \cdot \nabla f_{\Sigma}(Z_n)]$. In this case,

$$\begin{aligned} & \mathbb{E}[Z_n \cdot \nabla f_{\Sigma}(Z_n)] \\ &= \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} \mathbb{E}[(g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}])) \cdot \nabla f_{\Sigma}(Z_n)] \end{aligned}$$

Key observation

$$0 = \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} \mathbb{E}[(g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}])) \cdot \nabla f_{\Sigma^i}(\dot{Z}_n^i)],$$

where \dot{Z}_n^i and Σ^i are like Z_n and Σ , but removing the part of X in $[(i-1)/n, i/n]$

So what if $F = Z_n$?

Now we can write

$$\begin{aligned}\mathbb{E}[\mathbf{Z}_n \cdot \nabla f_{\Sigma}(\mathbf{Z}_n)] &= \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} \mathbb{E}[(g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}])) \\ &\quad \cdot \nabla(f_{\Sigma}(\mathbf{Z}_n) - f_{\Sigma^i}(\dot{\mathbf{Z}}_n^i))]\end{aligned}$$

New bottleneck:

So what if $F = Z_n$?

Now we can write

$$\begin{aligned}\mathbb{E}[\mathbf{Z}_n \cdot \nabla f_{\Sigma}(\mathbf{Z}_n)] &= \frac{1}{\sqrt{n}} \sum_{i=1}^{nt} \mathbb{E}[(g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) - \mathbb{E}[g(X_{\frac{i-1}{n}}, \mathcal{I}_{i,n}) \mid \mathcal{F}_{i-1}])) \\ &\quad \cdot \nabla(f_{\Sigma}(\mathbf{Z}_n) - f_{\Sigma^i}(\dot{\mathbf{Z}}_n^i))]\end{aligned}$$

New bottleneck: understand

$$f_{\Sigma}(\mathbf{Z}_n) - f_{\Sigma^i}(\dot{\mathbf{Z}}_n^i)$$

by means of Taylor approximations.

The mixed-Gaussian case

An easy criterion for mixed Gaussian convergence:

The mixed-Gaussian case

An easy criterion for mixed Gaussian convergence:

Theorem (Amorino, Jaramillo, Podolskij)

Consider a sequence \mathbf{F}_n , which is \mathcal{G} -measurable for some σ -algebra. If the convergence

$$\mathbb{E}[Y(\mathbf{F}_n) \cdot \nabla f(\mathbf{F}_n) - \text{Tr}[\text{Hess}[f](\mathbf{F}_n)\Sigma]] \rightarrow 0,$$

holds for all bounded \mathcal{G} -measurable Y and adequate test functions $h \in C^2(\mathbb{R}^r; \mathbb{R})$ then we obtain

$$S_n \xrightarrow{\text{Law}} \Sigma^{1/2} N,$$

where $N \sim \mathcal{N}_r(0, \text{id})$ is a standard r -dimensional normal variable defined on an extended space and independent of \mathcal{G} .

Main results

Let $X^{(1-m)}, \dots, X^{(2m)}$ be independent copies of \mathbf{X} , and define

$$\mathfrak{g}_n(\mathbf{x}) := \sum_{j=-m}^m \text{Cov}[g(\mathbf{x}, a_n X_{1/n}^{(1)}, \dots, a_n X_{1/n}^{(m+1)}), \\ g(\mathbf{x}, a_n X_{1/n}^{(j+1)}, \dots, a_n X_{1/n}^{(j+m+1)})].$$

Main results

Let $X^{(1-m)}, \dots, X^{(2m)}$ be independent copies of \mathbf{X} , and define

$$\mathfrak{g}_n(\mathbf{x}) := \sum_{j=-m}^m \text{Cov}[g(\mathbf{x}, a_n X_{1/n}^{(1)}, \dots, a_n X_{1/n}^{(m+1)}), \\ g(\mathbf{x}, a_n X_{1/n}^{(j+1)}, \dots, a_n X_{1/n}^{(j+m+1)})].$$

Theorem (Amorino, Jaramillo, Podolskij, 2023)

If \mathfrak{g}_n converges and $\alpha \in (0, 1)$, then

$$Z^{(n)} \xrightarrow{\text{Law}} \left\{ \int_0^t \sqrt{\lim_k \mathfrak{g}_k(\mathbf{X}_s)} W(ds) \right\}_{t \geq 0},$$

where W is a Brownian motion independent of \mathbf{X} .

Suppose that X is symmetric α -stable (including $\alpha = 2$), and define

$$d(\mu, \nu) := \sup_h \left| \int h d\mu - \int h d\nu \right|,$$

where the supremum is taken over all functions h satisfying $\|h^{(i)}\|_\infty \leq 1$ for $i = 0, 1, 2, 3$.

Main results (part II)

Theorem (Amorino, Jaramillo, Podolskij, 2023)

Given a fixed t , there exists a constant $C > 0$ depending only on g , such that:

- If $\alpha \in (1, 2)$,

$$d \left(Z_t^{(n)}, \int_0^t \sqrt{\lim_k g_k(X_s)} W(ds) \right) \leq Cn^{\frac{1}{2} - \frac{1}{\alpha}}.$$

- If $\alpha = 1$,

$$d \left(Z_t^{(n)}, \int_0^t \sqrt{\lim_k g_k(X_s)} W(ds) \right) \leq Cn^{-\frac{1}{2}} \log(n).$$

- If $\alpha \in (0, 1)$,

$$d \left(Z_t^{(n)}, \int_0^t \sqrt{\lim_k g_k(X_s)} W(ds) \right) \leq Cn^{-\frac{1}{2}}.$$

Theorem (Amorino, Jaramillo, Podolskij, 2023)

If g is symmetric and t is fixed, then for all $\alpha \in (1, 2]$, we have

$$d \left(Z_t^{(n)}, \int_0^t \sqrt{\lim_k g_k(X_s)} W(ds) \right) \leq Cn^{-\frac{1}{2}},$$

even for $\alpha = 2$.

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?
- Consider the case when X is a diffusion.

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?
- Consider the case when X is a diffusion.
- Remove independence of increments and consider:
 - (i) Gaussianity, replacing independence by weak dependence.
 - (ii) Replace independence with exchangeability properties.

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?
- Consider the case when X is a diffusion.
- Remove independence of increments and consider:
 - (i) Gaussianity, replacing independence by weak dependence.
 - (ii) Replace independence with exchangeability properties.
- Higher-order approximations (Edgeworth expansions).

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?
- Consider the case when X is a diffusion.
- Remove independence of increments and consider:
 - (i) Gaussianity, replacing independence by weak dependence.
 - (ii) Replace independence with exchangeability properties.
- Higher-order approximations (Edgeworth expansions).
- Allow more flexibility on g (e.g., $g(x) := \delta_0(x)$).

Future work

The structure of independent and stationary increments is convenient but not essential.

- Move from high-frequency observations to spaced observations (observe X_k instead of $X_{k/n}$).
- Remove stationarity of increments of X . Do we obtain integrals with time-changed Brownian motion?
- Consider the case when X is a diffusion.
- Remove independence of increments and consider:
 - (i) Gaussianity, replacing independence by weak dependence.
 - (ii) Replace independence with exchangeability properties.
- Higher-order approximations (Edgeworth expansions).
- Allow more flexibility on g (e.g., $g(x) := \delta_0(x)$).
- Understand the role of regularity of g (comparison with Itô integration).

Gracias!

Contacto

Arturo Jaramillo

jagil@cimat.mx



Chiara Amorino, Arturo Jaramillo, Mark Podolskij. Quantitative and stable limits of high-frequency statistics of Levy processes: a Stein's method approach.